Lecture 12: Learning Theory

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Intro. to Stats. Machine Learning COMP SCI 4401/7401



- Honours/Master/PhD projects
- eSELT (20 October to 7 November)
- Exam

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History

- Pionneered by Vapnik and Chervonenkis (1968, 1971), Sauer (1972), Shelah (1972) as Vapnik-Chevonenkis-Sauer Lemma
- Introduced in the west by Valiant (1984) under the name of "probably approximately correct" (PAC)

— with probability at least $1 - \delta$ (probably), any classifier from hypothesis class/set, if the classifier has low training error, it will have low generalisation error (approximately correct).

- Learnability and the VC dimension by Blumer *et al.* (1989), forms the basis of statistical learning theory
- Generalisation bounds, (1) SRM, Shawe-Taylor, Bartlett, Williamson, Anthony, (1998),
 - (2) Neural Networks, Bartlett (1998).
- Soft margin bounds, Cristianini, Shawe-Taylor (2000), Shawe-Taylor, Cristianini (2002)

History

- Apply Concentration inequalities, Boucheron *et al.* (2000), Bousquet, Elisseff (2001)
- Rademacher complexity, Koltchinskii, Panchenko (2000), Kondor, Lafferty (2002), Bartlett, Boucheron, Lugosi (2002), Bartlett, Mendelson (2002)
- PAC-Bayesian Bound proposed by McAllester (1999), improved by Seeger (2002) in Gaussian processes, applied to SVMs by Langford, Shawe-Taylor (2002), Tutorial by Langford (2005), greatly simplified proof by Germain *et al.* (2009).

Good books/tutorials

- J Shawe-Taylor, N Cristianini's book "Kernel Methods for Pattern Analysis", 2004
- V Vapnik's books "The nature of statistical learning theory", 1995 and "Statistical learning theory", 1998
- Online course "Learning from the Data", by Yaser Abu-Mostafa in Caltech.
- Bousquet *et al.*'s ML summer school tutorial "Introduction to Statistical Learning Theory", 2004

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Generalisation error Approximation error and estimation error Generalisation bounds

Generalisation error

 $\{(x_1, y_1), \cdots, (x_n, y_n) \sim P(X, Y)\}^1$, hypothesis function $g : \mathfrak{X} \to \mathfrak{Y}, \ \mathfrak{Y} = \{-1, 1\}.$

Generalisation error: error over all possible testing data from *P*, i.e. risk w.r.t. zero one loss $R(g) = \mathbb{E}_{(x,y)\sim P}[\mathbf{1}_{g(x)\neq y}]$.

Training error, i.e. empirical risk w.r.t. zero one loss $R_n(g) = \frac{1}{n} \sum_{i=1}^n [\mathbf{1}_{g(x_i) \neq y_i}].$

¹To simplify notation and make the results more general, we don't use boldface to distinguish vectors and scalers *i.e.* x, y, w can be vectors too.

Generalisation error Approximation error and estimation error Generalisation bounds

Approximation error and estimation error



Generalisation error Approximation error and estimation error Generalisation bounds

Generalisation bounds

$$g_{bayes} = \operatorname*{argmin}_{g} R(g)$$
$$g^* = \operatorname*{argmin}_{g \in \mathfrak{G}} R(g)$$
$$g_n = \operatorname*{argmin}_{g \in \mathfrak{G}} R_n(g)$$

Generalisation bounds:

$$egin{aligned} R(g_n) &\leq R_n(g_n) + B_1(n, \mathfrak{G}), & (1) \ R(g_n) &\leq R(g^*) + B_2(n, \mathfrak{G}), & (2) \ R(g_n) &\leq R(g_{bayes}) + B_3(n, \mathfrak{G}), & (3) \end{aligned}$$

where $B(n, \mathcal{G}) \ge 0$, and usually $B(n, \mathcal{G}) \to 0$ as $n \to +\infty$.



Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Capacity measure

Capacity/complexity of $\mathfrak{G} \downarrow \Rightarrow B(n, \mathfrak{G}) \downarrow$

How to measure the capacity/complexity of $\ensuremath{\mathfrak{G}}\xspace?$

• Counting the hypotheses in \mathcal{G} , *i.e.* $|\mathcal{G}|$.

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

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Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

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- Counting all possible outputs of the hypotheses
- Ability to fit noise

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Capacity measure

Capacity/complexity of $\mathfrak{G} \downarrow \Rightarrow B(n, \mathfrak{G}) \downarrow$

How to measure the capacity/complexity of \mathfrak{G} ?

- Counting the hypotheses in \mathcal{G} , *i.e.* $|\mathcal{G}|$.
- Counting all possible outputs of the hypotheses
- Ability to fit noise
- Divergence of the prior and posterior distributions (over classifiers)

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Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Counting the hypotheses

(a.k.a Hoeffding's inequality bound) For training examples $\{(x_1, y_1), \dots, (x_n, y_n)\}$, for a finite hypothesis set $\mathcal{G} = \{g_1, \dots, g_N\}$, for any $\delta \in (0, 1)$, with probability at least $1 - \delta$, $\forall g \in \mathcal{G}, R(g) \leq R_n(g) + \sqrt{\frac{\log N + \log(\frac{1}{\delta})}{2n}}$

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Proof (1)- Hoeffding's inequality

Theorem (Hoeffding)

Let Z_1, \dots, Z_n be n i.i.d. random variables with $f(Z) \in [a, b]$. Then for all $\epsilon > 0$, we have

$$\Pr\left(\left|\frac{1}{n}\sum_{i=1}^{n}f(Z_{i})-\mathbb{E}[f(Z)]\right|>\epsilon\right)\leq 2\exp\left(-\frac{2n\epsilon^{2}}{(b-a)^{2}}\right)$$

Let Z = (X, Y) and $f(Z) = \mathbf{1}_{g(X) \neq Y}$, we have $R(g) = \mathbb{E}(f(Z)) = \mathbb{E}_{(X,Y) \sim P}[\mathbf{1}_{g(X) \neq Y}]$ $R_n(g) = \frac{1}{n} \sum_{i=1}^n f(Z_i) = \frac{1}{n} \sum_{i=1}^n \mathbf{1}_{g(X_i) \neq Y_i}$ b = 1, a = 0 $\Rightarrow \Pr(|R(g) - R_n(g)| > \epsilon) \le 2 \exp(-2n\epsilon^2)$

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Proof (2) – for a hypothesis

$$\Pr(|R(g) - R_n(g)| > \epsilon) \le 2 \exp(-2n\epsilon^2)$$

Let
$$\delta = 2 \exp(-2n\epsilon^2) \Rightarrow \epsilon = \sqrt{\log(2/\delta)/2n}$$
.

⇒ For training examples $\{(x_1, y_1), \dots, (x_n, y_n)\}$, and for a hypothesis g, for any $\delta \in (0, 1)$ with probability at least $1 - \delta$,

$$R(g) \leq R_n(g) + \sqrt{rac{\log(rac{2}{\delta})}{2n}}$$

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Proof (3) – over finite many hypotheses

Let consider a finite hypothesis set $\mathcal{G} = \{g_1, \cdots, g_N\}$. Union bound

$$\Pr(\bigcup_{i=1}^{N} A_i) \leq \sum_{i=1}^{N} \Pr(A_i)$$

$$\Pr(|R(g) - R_n(g)| > \epsilon) \le 2 \exp(-2n\epsilon^2) \Rightarrow$$
$$\Pr(\exists g \in \mathfrak{G} : |R(g) - R_n(g)| > \epsilon) \le \sum_{i=1}^{N} \Pr(|R(g_i) - R_n(g_i)| > \epsilon)$$
$$\le 2N \exp(-2n\epsilon^2)$$

Let $\delta = 2N \exp(-2n\epsilon^2)$, we have, for any $\delta \in (0, 1)$, with probability at least $1 - \delta$,

$$\forall g \in \mathfrak{G}, R(g) \leq R_n(g) + \sqrt{\frac{\log N + \log(\frac{1}{\delta})}{2n}}$$

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Counting outputs

What if there are infinite many hypotheses $N = \infty$?

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Counting outputs

What if there are infinite many hypotheses $N = \infty$?

$$\sqrt{\frac{\log N + \log(\frac{1}{\delta})}{2n}} = \infty$$

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Counting outputs

What if there are infinite many hypotheses $N = \infty$?

Observation:

- For any x, only two possible outputs ($g(x) \in \{-1, +1\}$);
- 2 For any *n* training data at most 2^n different outputs of g(x).

What matters is the "expressive power" (Blumer *et al.* 1986,1989)(*e.g.* the number of different prediction outputs), not the cardinality of \mathcal{G} .

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Break

Take a break ...

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Growth function

Definition (Growth function)

The growth function (a.k.a Shatter coefficient) of \mathcal{F} with *n* points is

$$S_{\mathfrak{F}}(n) = \sup_{(z_1,\dots,z_n)} \left| \left\{ \left(f(z_1),\dots,f(z_n) \right) \right\}_{f\in\mathfrak{F}} \right|$$

i.e. maximum number of ways that n points can be classified by the hypothesis set \mathcal{F} .

Note: g can be a f, and \mathcal{G} can be a \mathcal{F} .

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Growth function

If no restriction on g, we know

$$\sup_{(z_1, z_2, z_3)} \left| \left\{ \left(g(z_1), g(z_2), g(z_3) \right) \right\} \right| = 2^3$$

When we restrict $g \in \mathcal{G}$,

$$S_{\mathfrak{G}}(3) = \sup_{(z_1, z_2, z_3)} \left| \left\{ \left(g(z_1), g(z_2), g(z_3) \right) \right\}_{g \in \mathfrak{G}} \right|.$$

i.e. counting all possible outputs that \mathcal{G} can express.

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Growth function

The growth function $S_{\mathcal{G}}(3) = 8$, if \mathcal{G} is the set of linear decision functions shown in the image below².



²The image is from http://www.svms.org/vc-dimension/

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Growth function

How about $S_{\mathcal{G}}(4)$?

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Growth function

How about $S_{\mathcal{G}}(4)$?



One g can not classify 4 points above correctly (two gs or a curve needed), which means $S_{g}(4) < 2^{4}$.

Picture courtesy of wikipedia

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VC dimension (1)

Definition (VC dimension)

The VC dimension (often denoted as h) of a hypothesis set G, is the largest n such that

$$S_{\mathfrak{G}}(n)=2^n.$$

h = 3 for \mathcal{G} being the set of linear decision functions in 2-D.

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VC dimension (2)

Lemma

Let \mathcal{G} be a set of functions with finite VC dimension h. Then for all $n \in \mathbb{N}$,

$$S_{\mathfrak{G}}(n) \leq \sum_{i=0}^{h} \binom{n}{i},$$

and for all $n \ge h$,

$$S_{\mathfrak{G}}(n) \leq (\frac{en}{h})^h.$$

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VC dimension (3)

Theorem (Growth function bound)

For any $\delta \in (0,1)$, with probability at least $1-\delta$, $\forall g \in \mathfrak{G}$

$$R(g) \leq R_n(g) + 2\sqrt{2\frac{\log S_{\mathcal{G}}(2n) + \log(\frac{2}{\delta})}{n}}$$

Thus for all $n \ge h$, since $S_{\mathfrak{G}}(n) \le (\frac{en}{h})^h$, we have

Theorem (VC bound)

For any $\delta \in (0,1)$, with probability at least $1-\delta$, $\forall g \in \mathfrak{G}$

$$R(g) \leq R_n(g) + 2\sqrt{2rac{h\lograc{2en}{h} + \log(rac{2}{\delta})}{n}}.$$

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VC dimension (4)

Assume $x \in \mathbb{R}^d$, $\Phi(x) \in \mathbb{R}^D$ (Note D can be $+\infty$).

• linear
$$\langle x, w \rangle$$
, $h = d + 1$

- polynomial $(\langle x,w
 angle+1)^p$, $h={d+p-1\choose p}+1$
- Gaussian RBF exp $\left(-\frac{\|\mathbf{x}-\mathbf{x}'\|^2}{\sigma^2}\right)$, $h = +\infty$.
- Margin γ, h ≤ min{D, [4R²/γ²]}, where the radius
 R² = maxⁿ_{i=1} ⟨Φ(x_i), Φ(x_i)⟩ (assuming data are already centered)

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

Ability to fit noise

Definition (Rademacher complexity)

Given $S = \{z_1, \dots, z_n\}$ from a distribution P and a set of real-valued functions \mathcal{G} , the empirical Rademacher complexity of \mathcal{G} is the random variable

$$\hat{\mathcal{R}}_n(\mathcal{G},S) = \mathbb{E}_{\boldsymbol{\sigma}} \Big[\sup_{\boldsymbol{g} \in \mathcal{G}} \Big| \frac{2}{n} \sum_{i=1}^n \sigma_i \boldsymbol{g}(z_i) \Big| \Big],$$

where $\sigma = \{\sigma_1, \cdots, \sigma_n\}$ are independent uniform $\{\pm 1\}$ -valued (Rademacher) random variables. The Rademacher complexity of \mathcal{G} is

$$\mathfrak{R}_n(\mathfrak{G}) = \mathbb{E}_{\mathcal{S}}[\hat{\mathfrak{R}}_n(\mathfrak{G}, \mathcal{S})] = \mathbb{E}_{\mathcal{S}\sigma}\left[\sup_{g \in \mathfrak{G}} \left|\frac{2}{n}\sum_{i=1}^n \sigma_i g(z_i)\right|\right]$$

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First sight

$$\sup_{g\in\mathfrak{G}}\left|\frac{2}{n}\sum_{i=1}^{n}\sigma_{i}g(z_{i})\right|$$

- measures the best correlation between $g \in \mathcal{G}$ and random label (*i.e.* noise) $\sigma_i \sim U(\{-1, +1\})$.
- ability of \mathcal{G} to fit noise.
- the smaller, the less chance of detected pattern being spurious

• if
$$|\mathcal{G}| = 1$$
, $\mathbb{E}_{\sigma}\left[\sup_{g \in \mathcal{G}} \left|\frac{2}{n} \sum_{i=1}^{n} \sigma_{i}g(z_{i})\right|\right] = 0$.

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Rademacher bound

Theorem (Rademacher)

Fix $\delta \in (0,1)$ and let \mathfrak{G} be a set of functions mapping from Z to [a, a+1]. Let $S = \{z_i\}_{i=1}^n$ be drawn i.i.d. from P. Then with probability at least $1 - \delta$, $\forall g \in \mathfrak{G}$,

$$\mathbb{E}_{P}[g(z)] \leq \hat{\mathbb{E}}[g(z)] + \mathcal{R}_{n}(\mathfrak{G}) + \sqrt{\frac{\ln(2/\delta)}{2n}} \\ \leq \hat{\mathbb{E}}[g(z)] + \hat{\mathcal{R}}_{n}(\mathfrak{G}, S) + 3\sqrt{\frac{\ln(2/\delta)}{2n}},$$

where $\hat{\mathbb{E}}[g(z)] = \frac{1}{n} \sum_{i=1}^{n} g(z_i)$

Note: $\hat{\mathcal{R}}_n(\mathcal{G}, S)$ is computable whereas $\mathcal{R}_n(\mathcal{G})$ is not.

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Properties of empirical Rademacher complexity

Let $\mathcal{F}, \mathcal{F}_1, \dots, \mathcal{F}_m$ and \mathcal{G} be classes of real functions. Let $S = \{z_i\}_{i=1}^n$ i.i.d. from any unknown but fixed P.Then **1** If $\mathcal{F} \subseteq \mathcal{G}$, then $\hat{\mathcal{R}}_n(\mathcal{F}, S) \leq \hat{\mathcal{R}}_n(\mathcal{G}, S)$ 2 For every $c \in \mathbb{R}$, $\hat{\mathcal{R}}_n(c\mathcal{F}, S) = |c|\hat{\mathcal{R}}_n(\mathcal{F}, S)$ 3 $\hat{\mathcal{R}}_n(\sum_{i=1}^m \mathcal{F}_i, S) < \sum_{i=1}^m \hat{\mathcal{R}}_n(\mathcal{F}_i, S)$ • For any function h, $\hat{\mathcal{R}}_n(\mathcal{F}+h, S) \leq \hat{\mathcal{R}}_n(\mathcal{F}, S) + 2\sqrt{\hat{\mathbb{E}}[h^2]/n}$ $\hat{\mathfrak{R}}_n(\mathfrak{F},S) = \hat{\mathfrak{R}}_n(\operatorname{conv}(\mathfrak{F}),S)$ **o** If $\mathcal{A} : \mathbb{R} \to \mathbb{R}$ is Lpschitz with constant L > 0 (*i.e.* $|\mathcal{A}(a) - \mathcal{A}(a')| \leq L|a - a'|$ for all $a, a' \in \mathbb{R}$), and $\mathcal{A}(0) = 0$, then $\hat{\mathcal{R}}_n(\mathcal{A} \circ \mathcal{F}, S) < 2L\hat{\mathcal{R}}_n(\mathcal{F}, S)$

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An example

Let $S = \{(x_i, y_i)\}_{i=1}^n \sim P^n$. $y_i \in \{-1, +1\}$ One form of soft margin binary SVMs is

$$\begin{split} \min_{\substack{w,\gamma,\xi}} &-\gamma + C \sum_{i=1}^{n} \xi_{i} \\ \text{s.t. } y_{i} \left\langle \phi(x_{i}), w \right\rangle \geq \gamma - \xi_{i}, \xi_{i} \geq 0, \|w\|^{2} = 1 \end{split}$$
(4)

- The Rademacher Margin bound (next slide) applies.
- $\hat{\mathcal{R}}_n(\mathcal{G}, S)$ is essential, where $\mathcal{G} = \{-yf(x; w), f(x; w) = \langle \phi(x_i), w \rangle, \|w\|^2 = 1\}.$

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Rademacher Margin bound

Theorem (Margin)

Fix $\gamma > 0, \delta \in (0, 1)$, let \mathcal{G} be the class of functions mapping from $\mathfrak{X} \times \mathfrak{Y} \to \mathbb{R}$ given by g(x, y) = -yf(x), where f is a linear function in a kernel-defined feature space with norm at most 1. Let $S = \{(x_i, y_i)\}_{i=1}^n$ be drawn i.i.d. from P(X, Y) and let $\xi_i = (\gamma - y_i f(x_i))_+$. Then with probability at least $1 - \delta$ over sample of size n, we have

$$\mathbb{E}_{\mathcal{P}}[\mathbf{1}_{y\neq \operatorname{sgn}(f(x))}] \leq \frac{1}{n\gamma} \sum_{i=1}^{n} \xi_{i} + \frac{4}{n\gamma} \sqrt{\operatorname{tr}(\mathbf{K})} + 3\sqrt{\frac{\ln(2/\delta)}{2n}},$$

- data dependency come through training error and margin
- tighter than VC bound $(\frac{4}{n\gamma}\sqrt{tr(\mathbf{K})} \le \frac{4}{n\gamma}\sqrt{nR^2} \le 4\sqrt{\frac{R^2}{n\gamma^2}})$

History Generalisation bounds Capacity measure Divergence of the prior and posterior distributions

PAC-bayes bounds

Assume \hat{Q} is the prior distribution over classifiers $g \in \mathcal{G}$ and Q is any (could be the posterior) distribution over the classifiers.

PAC-bayes bounds on:

- Gibbs classifier: $G_Q(x) = g(x), g \sim Q$ risk: $R(G_Q) = \mathbb{E}_{(x,y) \sim P, g \sim Q}[\mathbf{1}_{g(x) \neq y}]$ (McAllester98,99,01,Germain *et al.* 09)
- Average classifier: $B_Q(x) = \operatorname{sgn}[\mathbb{E}_{g \sim Q} g(x)]$ risk: $R(B_Q) = \mathbb{E}_{(x,y) \sim P}[\mathbf{1}_{\mathbb{E}_Q[g(x)] \neq y}]$ (Langford01, Zhu&Xing09)
- Single classifier: $g \in \mathcal{G}$. risk: $R(g) = \mathbb{E}_{(x,y)\sim P}[\mathbf{1}_{g(x)\neq y}]$ (Langford01,McAllester07)

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Why bounding Gibbs classifier is enough?

PAC-bayes bounds on Gibbs classifier yield the bounds on average classifier and single classifier.

 $R(G_Q)$ (original PAC-Bayes bounds)

 $\Downarrow :: R(B_Q)/2 \le R(G_Q)$

 $R(B_Q)$ (PAC-Bayes margin bound for boostings)

 \Downarrow via picking a "good" prior \hat{Q} and posterior Q over g

R(g) (PAC-Bayes margin bound for SVMs)

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PAC-Bayesian bound on Gibbs Classifier (1)

Theorem (Gibbs (McAllester99,03))

For any distribution P, for any set \mathfrak{G} of the classifiers, any prior distribution \hat{Q} of \mathfrak{G} , any $\delta \in (0, 1]$, we have

$$\Pr_{S \sim P^n} \left\{ \forall Q \text{ on } \mathcal{G} : R(G_Q) \leq R_S(G_Q) + \sqrt{\frac{1}{2n-1} \left[\mathcal{K}L(Q || \hat{Q}) + \ln \frac{1}{\delta} + \ln n + 2 \right]} \right\} \geq 1 - \delta.$$

where $KL(Q||\hat{Q}) = \mathbb{E}_{g \sim Q} \ln \frac{Q(g)}{\hat{Q}(g)}$ is the KL divergence.

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PAC-Bayesian bound on Gibbs Classifier (2)

Theorem (Gibbs (Seeger02 and Langford05))

For any distribution P, for any set \mathfrak{G} of the classifiers, any prior distribution \hat{Q} of \mathfrak{G} , any $\delta \in (0, 1]$, we have

$$\Pr_{S\sim P^n} \left\{ \forall Q \text{ on } \mathcal{G} : kl(R_S(G_Q), R(G_Q)) \leq \frac{1}{n} \left[KL(Q||\hat{Q}) + \ln \frac{n+1}{\delta} \right] \right\} \geq 1 - \delta.$$

where

$$kl(q,p)=q\lnrac{q}{p}+(1-q)\lnrac{1-q}{1-p}.$$

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PAC-Bayesian bound on Gibbs Classifier (3)

Since

$$kl(q,p) \geq (q-p)^2,$$

The theorem Gibbs (Seeger02 and Langford05) yields

$$\Pr_{S \sim P^n} \left\{ \forall Q \text{ on } \mathcal{G} : R(G_Q) \right\} \le R_S(G_Q) + \sqrt{\frac{1}{n} \left[\mathsf{KL}(Q || \hat{Q}) + \ln \frac{n+1}{\delta} \right]} \right\} \ge 1 - \delta.$$

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PAC-Bayesian bound on Average Classifier

Theorem (Average (Langford *et al.* 01))

For any distribution P, for any set \mathfrak{G} of the classifiers, any prior distribution \hat{Q} of \mathfrak{G} , any $\delta \in (0, 1]$, and any $\gamma > 0$, we have

$$\Pr_{S \sim P^n} \left\{ \forall Q \text{ on } \mathcal{G} : R(B_Q) \leq \Pr_{(\mathbf{x}, y) \sim S} \left(y \mathbb{E}_{g \sim Q}[g(x)] \leq \gamma \right) \right. \\ \left. + O\left(\sqrt{\frac{\gamma^{-2} \mathcal{K} \mathcal{L}(Q || \hat{Q}) \ln n + \ln n + \ln \frac{1}{\delta}}{n}} \right) \right\} \geq 1 - \delta.$$

Zhu& Xing09 extended to structured output case.

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PAC-Bayesian bound on Single Classifier

Assume $g(x) = \langle w, \phi(x) \rangle$ and rewrite R(g) as R(w).

Theorem (Single (McAllester07))

For any distribution P, for any set \mathfrak{G} of the classifiers, any prior distribution \hat{Q} over w, any $\delta \in (0, 1]$, and any $\gamma > 0$, we have

$$\Pr_{S \sim P^n} \left\{ \forall w \sim \mathcal{W} : R(w) \leq \Pr_{(\mathbf{x}, y) \sim S} \left(y \langle w, \phi(x) \rangle \leq \gamma \right) \right. \\ \left. + O\left(\sqrt{\frac{\gamma^{-2} \frac{\|w\|^2}{2} \ln(n |\mathcal{Y}|) + \ln n + \ln \frac{1}{\delta}}{n}} \right) \right\} \geq 1 - \delta.$$

Counting the hypotheses Counting outputs Ability to fit noise Divergence of the prior and posterior distributions

That's all

Good luck with the exam.